#### Soc504: Basics

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 $<sup>^1 \</sup>rm{These}$  slides are a very lightly remixed version of slides for Gary King's course Gov2001 at Harvard.

### Followup

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• Questions?

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- Questions?
- Replication Stories?











#### 2 Useful Distributions

3 Concluding Thoughts

Appendix: More Probability Distributions

Solve probability problems

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- 2 Evaluate estimators

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- 2 Evaluate estimators
- Salculate features of probability densities
- Transform statistical results into quantities of interest
- Get the right answer: students get the right answer far more frequently by using simulation than math

Survey Sampling

Simulation

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1. Learn about a population	1. Learn about a distribu-
by taking a random sample	tion by taking random draws
from it	from it

Survey Sampling	Simulation
1. Learn about a population	1. Learn about a distribu-
by taking a random sample	tion by taking random draws
from it 2. Use the random sample to estimate a feature of the population	from it 2. Use the random draws to approximate a feature of the distribution

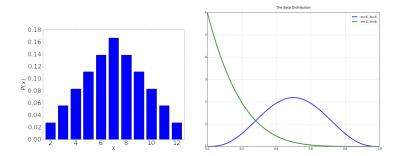
Survey Sampling	Simulation
1. Learn about a population	1. Learn about a distribu-
by taking a random sample	tion by taking random draws
from it 2. Use the random sample	from it 2. Use the random draws to
to estimate a feature of the	approximate a feature of the
population	distribution
3. The estimate is arbitrarily	3. The approximation is ar-
precise for large <i>n</i>	bitrarily precise for large $M$

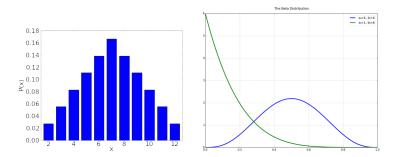
Survey Sampling	Simulation
1. Learn about a population	1. Learn about a distribu-
by taking a random sample	tion by taking random draws
from it 2. Use the random sample to estimate a feature of the population 3. The estimate is arbitrarily precise for large <i>n</i> 4. Example: estimate the mean of the population	<ul> <li>from it</li> <li>2. Use the random draws to approximate a feature of the distribution</li> <li>3. The approximation is arbitrarily precise for large M</li> <li>4. Example: Approximate the mean of the distribution</li> </ul>

```
sims <- 1000
WinNoSwitch <- 0
WinSwitch <- 0
doors <- c(1, 2, 3)</pre>
```

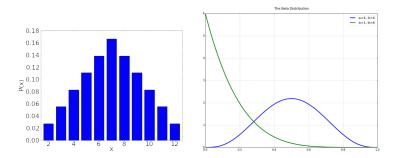
```
sims <- 1000
  WinNoSwitch <- 0
  WinSwitch <-0
  doors <- c(1, 2, 3)
  for (i in 1:sims) {
    WinDoor <- sample(doors, 1)</pre>
    choice <- sample(doors, 1)</pre>
    if (WinDoor == choice)
                                                   # no switch
       WinNoSwitch <- WinNoSwitch + 1
    doorsLeft <- doors[doors != choice]</pre>
                                                   # switch
    if (any(doorsLeft == WinDoor))
       WinSwitch <- WinSwitch + 1
 }
cat("Prob(Car | no switch)=". WinNoSwitch/sims. "\n")
cat("Prob(Car | switch)=", WinSwitch/sims, "\n")
```

#### Pr(car|No Switch) Pr(car|Switch) .324 .676 .345 .655 .320 .680 .327 .673

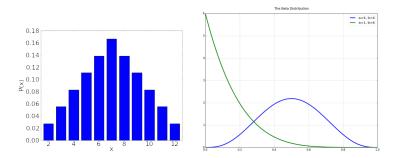




Required for a PDF: Denoted:  $\mathbb{P}(y)$ 



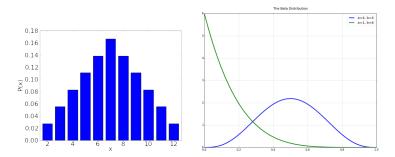
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 for every  $y$ 

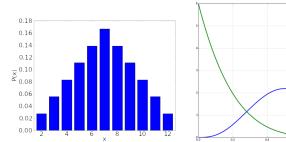
$$\bigcirc \sum_{y} \mathbb{P}(y) = 1$$



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•  $\mathbb{P}(y) \ge 0$  for every y

2 
$$\sum_{y} \mathbb{P}(y) = 1$$
 or  
 $\int_{-\infty}^{\infty} \mathbb{P}(y) dy = 1$ 



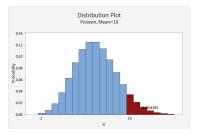
a=4, b=4 - a=1, b=6

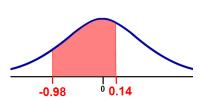
The Beta Distribution

Required for a PDF: Denoted:  $\mathbb{P}(y)$ 

•  $\mathbb{P}(y) \geq 0$  for every y

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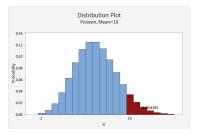


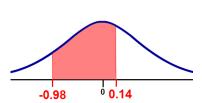
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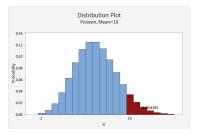
# Required for a PDF: Denoted: $\mathbb{P}(u)$

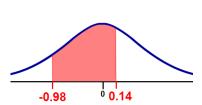
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- $\Pr(a \le Y \le b) = \int_a^b \mathbb{P}(y) dy$
- For discrete:  $\Pr(y) = \mathbb{P}(y)$
- For continuous: Pr(y) = 0

What you should know about every pdf

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• The assignment of a probability or probability density to every conceivable value of *Y<sub>i</sub>* 

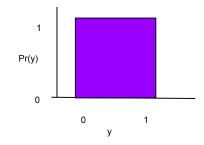
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- The first principles

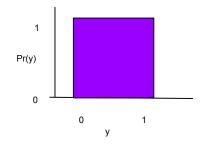
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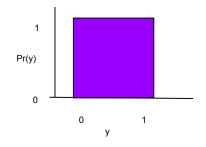
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- How to verify that the final expression is indeed a proper density



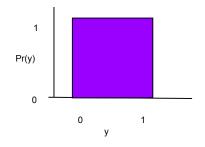


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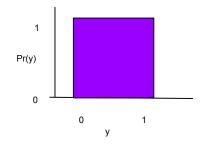
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 if  $a < b$ ,  $c < d$ , and  $b - a = d - c$ .

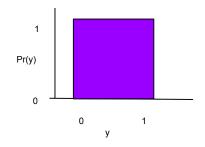


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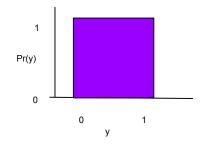
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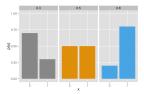


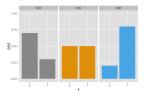
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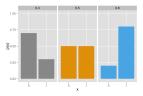
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- Is it a pdf? How do you know?
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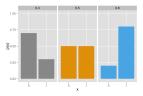




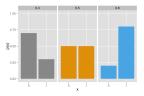
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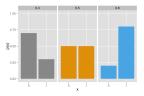
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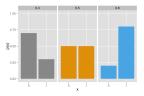
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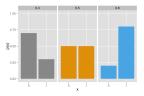


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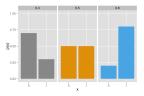


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  - $\Pr(Y_i = 1 | \pi_i) = \pi_i$ ,  $\Pr(Y_i = 0 | \pi_i) = 1 \pi_i$
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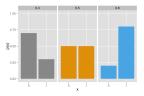


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$$\blacktriangleright \implies \mathsf{Pr}(Y_i = y | \pi_i) = \pi_i^y (1 - \pi_i)^{1 - y}$$



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- $\blacktriangleright \implies \mathsf{Pr}(Y_i = y | \pi_i) = \pi_i^y (1 \pi_i)^{1-y}$
- Alternative notation:  $Pr(Y_i = y | \pi_i) = Bernoulli(y | \pi_i) = f_b(y | \pi_i)$

$$E(Y) = \sum_{\text{all } y} y \mathbb{P}(y)$$

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• How do we compute  $E(Y^2)$ ?

# Expected values of functions of random variables

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$$E[g(Y)] = \sum_{\text{all } y} g(y) \mathbb{P}(y)$$

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#### $V(Y) = E[(Y - E(Y))^2]$ (The definition)

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Think about where the maximum is. Does it accord with your intuition?

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In R:

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sims <- 1000  # set parameters
bernpi <- 0.2
u <- runif(sims)  # uniform sims
y <- as.integer(u < bernpi)
y  # print results</pre>
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• In practice, can use rbinom(size=1)

First principles:

- *N* iid Bernoulli trials,  $y_1, \ldots, y_N$
- The trials are independent
- The trials are identically distributed
- We observe  $Y = \sum_{i=1}^{N} y_i$

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Explanation:

- $\binom{N}{y}$  because (1 0 1) and (1 1 0) are both y = 2.
- $\pi^{y}$  because y successes with  $\pi$  probability each (product taken due to independence)
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First principles:

- *N* iid Bernoulli trials, *y*<sub>1</sub>,...,*y*<sub>N</sub>
- The trials are independent
- The trials are identically distributed
- We observe  $Y = \sum_{i=1}^{N} y_i$

Density:

$$\mathbb{P}(Y = y|\pi) = \binom{N}{y} \pi^{y} (1-\pi)^{N-y}$$

Explanation:

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- $\pi^{y}$  because y successes with  $\pi$  probability each (product taken due to independence)
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- Mean  $E(Y) = N\pi$

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- What can you do with the simulations?

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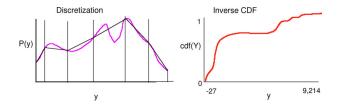
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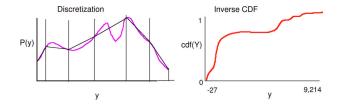
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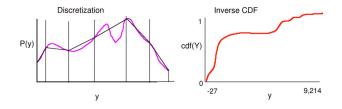
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- In the midrange future we might be using quantum computers for this.

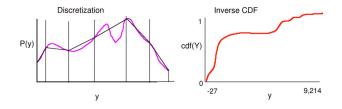




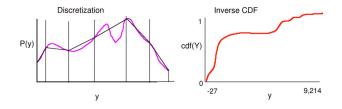
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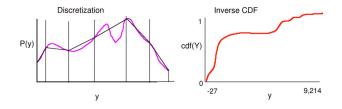
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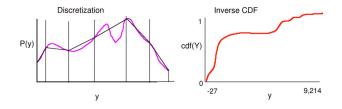
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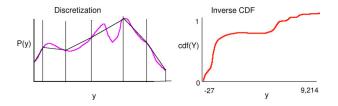
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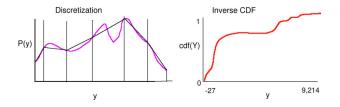


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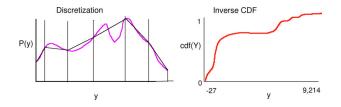


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- (Works for a few dimensions, but infeasible for many)



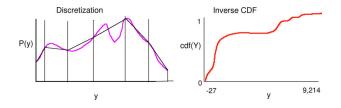


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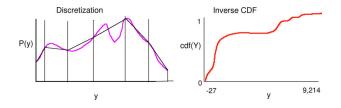


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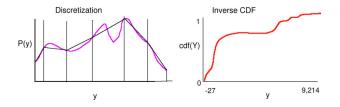
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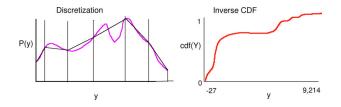


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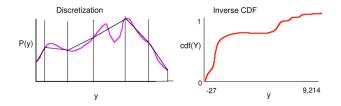


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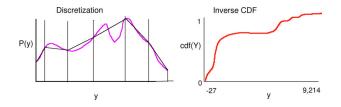




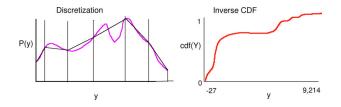
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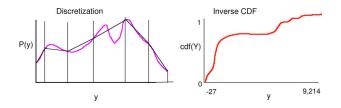
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- Also a decent literature on drawing samples with different speed/accuracy tradeoffs











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3 Concluding Thoughts



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• Simulating <u>once</u> from this density produces k numbers. Special algorithms are used to generate normal random variates (in R, mvrnorm(), from the MASS library).

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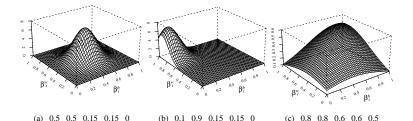
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# Truncated bivariate normal examples (for $\beta^b$ and $\beta^w$ )



Parameters are  $\mu_1$ ,  $\mu_2$ ,  $\sigma_1$ ,  $\sigma_2$ , and  $\rho$ .



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By Next Wednesday: Read UPM Chapter 4











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3 Concluding Thoughts



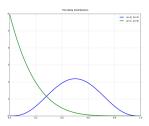
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Reparameterization like this will be key throughout the course.

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(First principles are easy to see from this too.)

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- Add up the  $\tilde{z}$ 's to get  $y = \sum_{j}^{N} \tilde{z}_{j}$ , which is a draw from the beta-binomial.

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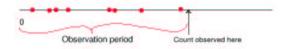
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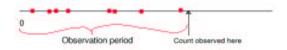
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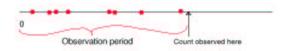
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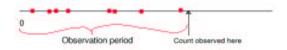
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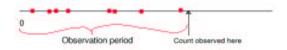
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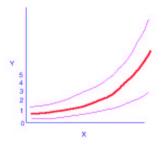
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- How to simulate? We'll use canned random number generators.

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$$= \frac{\Gamma\left(\frac{\phi}{\sigma^2 - 1} + y_i\right)}{y_i! \Gamma\left(\frac{\phi}{\sigma^2 - 1}\right)} \left(\frac{\sigma^2 - 1}{\sigma^2}\right)^{y_i} (\sigma^2)^{\frac{-\phi}{\sigma^2 - 1}}$$