Soc504: Inference

Brandon Stewart¹

Princeton

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 $^{^1 \}rm Much$ of the material in section 2-7 is edited from Gary King's slides for Gov2000 at Harvard. Section 8 is heavily influenced by Patrick Lam.

Where We've Been and Where We're Going...

- Last Week
 - Intro and Class Overview
- This Week+
 - Theories of inference
 - Likelihood Estimation
 - Simulation
- Next Week+
 - Generalized Linear Models
- Long Run
 - ▶ likelihood \rightarrow GLMs \rightarrow advanced methods



• Questions?

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- Note that today we will be discussing things that border on philosophy. Thus it is particularly important that you ask questions. Often the simplest questions are the most profound!



- 2 Likelihood Inference
- 3 Bayesian Inference
 - Neyman-Pearson
- 5 Likelihood Example
- 6 Properties and Tests

7 Simulation

8 Fun With Bayes



Likelihood Inference

- 3 Bayesian Inference
- 4 Neyman-Pearson
- 5 Likelihood Example
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- 7 Simulation
- 8 Fun With Bayes

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- We've been using algorithms of various sorts for a long time, least squares dates back to Legendre and Gauss 1795-1805.
- While an algorithm tells us what to compute and provides a summary of the data, inference answers the question of why we are doing something (i.e. what properties it has).
- For our purposes the central question of inference will be, how do we assess the accuracy of an estimate?

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- We often talk about this as frequentists posing the question: 'what would happen if we reran the same situation over and over again?'
- Why is this hard? Well we need to calculate properties of an estimator obtained from a true distribution *F* even though *F* is unknown.

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Figure 3.5 Bayesian inference proceeds vertically, given x; frequentist inference proceeds horizontally, given μ .

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Inference

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- We will also see that likelihood lends itself nicely to situations where we care a lot about the outcome rather than the coefficients themselves.
- For those interested Stigler's "The epic story of maximum likelihood" is a fantastic account of the history of the idea.

A Perspective on a Historical Arc (Efron and Hastie 2016)



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3. A more reasonable, limited goal. Let $M = \{M^*, \theta\}$, where M^* is assumed & θ is to be estimated:

$$\mathbb{P}(\theta|y, M^*) \equiv \mathbb{P}(\theta|y)$$

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- 8. The two differ on what is fixed and what is random



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History

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6. $L(\theta|y)$ is a function: for y fixed at the observed values, it gives the "likelihood" of any value of θ .

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- 11. The likelihood principle: the data only affect inferences through the likelihood function

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- Uncertainty of point estimate: curvature at the maximum

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- Notational note: log in math is almost always used as short-hand for the natural log (In) as opposed to the base-10 log.

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$$Y = (Y_1, Y_2, \dots, Y_n)$$

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For a fixed set of observations, what does this look like?

Example 1: Bernoulli Trials: Simulated Example



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Uncertainty About Mode $\pi^* = \overline{Y}$ maximizes $L(\pi | \mathbf{Y})$.

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Second derivative captures this curvature



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Interpretation 2: The Bayesian Theory of Inference

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• $\mathbb{P}(\theta)$, the prior density — the way Bayes differs from likelihood

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- 2. An opportunity: a way of getting other information outside the data set into the model
- 3. An annoyance: the "other information" is required
- 4. A philosophical assumption that nonsample information should matter (as it always does) and be formalized and included in all inferences.

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- 4. If variables A and B are both unknown, then the distribution of A alone is $P(A) = \int \mathbb{P}(A, B) dB = \int P(A|B)P(B) dB$.





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- Bayesian inference obeys the likelihood principle: the data set only affects inferences through the likelihood function
- If $\mathbb{P}(\theta) = 1$, i.e., is uniform in the relevant region, then $L(\theta|y) = \mathbb{P}(\theta|y)$.

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- A perspective of growing importance is empirical Bayes which we will discuss later in the semester.



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- 4. Methods for applied researchers: either useful or irrelevant \rightarrow learn something then validate it.

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- This motivates different views of the core material such as agnostic and robust statistics.



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$$\mathbb{P}(y|\mu) \equiv \mathbb{P}(y_1, \dots, y_n|\mu_1, \dots, \mu_n) = \prod_{i=1}^n f_{stn}(y_i|\mu_i)$$
$$= \prod_{i=1}^n (2\pi)^{-1/2} \exp\left(\frac{-(y_i - \mu_i)^2}{2}\right)$$

reparameterizing with $\mu_i = \beta$:

- 1. $Y_i \sim f_{stn}(y_i|\mu_i)$, normal stochastic component
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• What can you do with this probability density?

Stewart (Princeton)

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β



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- 5. No reason to summarize this curve with only the MLE

• The problem of Flatland



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• The problem of Flatland

• Graphs



- The problem of Flatland
- Graphs
- The curse of dimensionality



- The problem of Flatland
- Graphs
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- Maximum



- The problem of Flatland
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- The curvature at the maximum (standard errors, about which more shortly)



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- ▶ Not a sharp divide- some analytic work helps numerical optimization

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- To do this, we write down a probability model for the data.

Empirical distribution of wall-punching ages



Ages of ER patients who punched a wall in 2014

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- We assume that each Y_i ~ Log-Normal(μ, σ²), and that each Y_i is independently and identically distributed. (Later we could extend this model by adding covariates (e.g. μ_i = X_iβ)).

The density of the log-normal distribution is given by

$$f(Y_i|\mu,\sigma^2) = \frac{1}{Y_i\sigma\sqrt{2\pi}} \exp\left(-\frac{(\ln(Y_i)-\mu)^2}{2\sigma^2}\right)$$

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Basically the same as saying $ln(Y_i)$ is normally distributed!
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• Unfortunately, $f(\mathbf{Y}|\mu, \sigma^2)$ is an *n*-dimensional density, and *n* is huge! How do we simplify this? The *i.i.d.* assumption lets us factor the density!

$$\mathcal{L}(\mu, \sigma^2 | \mathbf{Y}) \propto \prod_{i=1}^N f(Y_i | \mu, \sigma^2)$$

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- This is why we typically work with the log-likelihood (often denoted ℓ). Because taking the log is a monotonic transformation, it retains the proportionality!

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Plotting the log-likelihood



Figure: Contour plot of the log-likelihood for different values of μ and σ

Plotting the likelihood



Figure: Plot of the log-likelihood for different values of μ and σ

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Stewart (Princetoni	

Plotting the likelihood



Conditional log-likelihood varying mu, setting sigma=2

Figure: Plot of the conditional log-likelihood of μ given $\sigma = 2$

• Example 1: $\mu = 4$, $\sigma = .2$: Log-likelihood = -18048.79

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- Example 2: $\mu = 3.099$, $\sigma = 0.379$: Log-likelihood = -4461.054 (actually the MLE)!
- Let's plot the implied distribution of Y_i for each parameter set over the empirical histogram!



Ages of ER patients who punched a wall in 2014

Figure: Empirical distribution of ages vs. log-normal with $\mu = 4$ and $\sigma = .2$



Figure: Empirical distribution of ages vs. log-normal using MLEs of parameters



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Invariance to sampling plans

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- Asymptotic efficiency. The MLE contains as much information as can be packed into a point estimator.

Sampling distributions of the MLE: CLT vs LLN



Stewart (Princeton





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where r is the observed value of R and m is the number of restricted parameters.





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- So only if r >> m will the test parameters be clearly different from zero.
- Disadvantage: Too many likelihood ratio tests may be required to test all points of interest
- Thus, it might be nice to have a summary of uncertainty for every parameter separately → standard errors





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- the more information exists in the MLE
- the larger the likelihood ratio would be in comparing the MLE with <u>any</u> other parameter value.

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8. We invert the curvature to provide a statistical interpretation:

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- 9. This is an estimate of a quadratic approximation to the log-likelihood.

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- In certain settings we can still prove the point estimate is consistent and derive consistent estimators of the sampling variance (heteroskedasticity and serial correlation in normal model, clustering in logit and probit models, overdispersion in GLMs)



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 - ▶ We'll discuss later how how to improve the approximation.

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 E_i The number of electoral college votes for each state in 2016
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- Compute and save $\hat{V}(\hat{ heta})$, which is k+2 imes k+2

• Mathematical Form:

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• An R function:

```
ll.normal <- function(par, X, Y) {
X <- as.matrix(cbind(1, X))
beta <- par[1:ncol(X)]
sigma2 <- exp(par[ncol(X) + 1])
-1/2 * sum( log(sigma2) + ((Y - X %*% beta)^2)/sigma2 )
}</pre>
```

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- Repeat Steps 1–3 M = 1,000 times, and plot a histogram of the results.

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 - Or, in our preferred notation, draw $\tilde{y}_{i,2016}$ from $N(X_{i,2016}\tilde{\beta},\tilde{\sigma}^2)$

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(calculated before the election by Gelman and King)

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• For what applications would this model be informative?

An Outline of the Research Process



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- We will talk about this more in the last couple of weeks.



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e.g. when rolling a die, the probability of a 3 or 5 is just $\frac{1}{6} + \frac{1}{6}$

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- "objective", dominant paradigm in statistics, and cheerleaders incl Fischer.

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- example: our prior over a coin's outcomes (Bernoulli process) might be $p = \frac{1}{2}$ or $p = \frac{1}{3}$ or p = 1 ('degenerate')— we can then conduct our trials (the tosses themselves). Alternatively, we might have a prior on the value of some $\hat{\beta}$ in a regression.

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B: Bayesian methods allow us to easily estimate models that are too hard to estimate (cannot computationally find the MLE) or unidentified (no unique MLE exists) with non-Bayesian methods. Bayesian methods also allow us to incorporate prior/qualitative information into the model.

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There is a Bayesian way to do any non-Bayesian parametric model.

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- The outcomes are independent.

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Bayesian priors are just adding pseudo observations to the data.

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Big Point: Bayesian inference necessitates the estimation of distributions rather than parameters

Uninformative Beta(1,1) Prior



Beta(2,12) Prior



Uninformative Beta(1,1) Prior (n=1000)



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$$Beta(y + \alpha, n - y + \beta) = \frac{Binomial(n, \pi) \times Beta(\alpha, \beta)}{p(y)}$$

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Conjugate models are great because we can find the exact posterior, but...

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$$\begin{split} \rho(\boldsymbol{\alpha},\boldsymbol{\beta},\boldsymbol{\theta},\boldsymbol{\sigma},\boldsymbol{\pi},\boldsymbol{\tau}|\boldsymbol{Y}) & \propto & \prod_{k=1}^{K} \prod_{s=1}^{S} \frac{\exp(-\frac{\alpha_{ks}}{1/4})}{1/4} \times \frac{\Gamma(\sum_{w=1}^{W} \lambda_w)}{\prod_{w=1}^{W} \Gamma(\lambda_w)} \prod_{w=1}^{W} \theta_{k,w}^{\lambda_w-1} \times \\ & \prod_{i=1}^{n} \prod_{t=2005}^{2007} \sum_{s=1}^{S} \left[\beta_s \frac{\Gamma(\sum_{k=1}^{K} \alpha_{ks})}{\prod_{k=1}^{K} \Gamma(\alpha_{ks})} \prod_{k=1}^{K} \pi_{itk}^{\alpha_{ks}-1} \prod_{j=1}^{D_{it}} \prod_{k=1}^{K} \left[\pi_{itk} \prod_{w=1}^{W} \theta_{kw}^{y_{ijtw}} \right]^{\tau_{ijtk}} \right]^{\sigma_{its}} \end{split}$$

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All Causes (f), n = 2 N 0 Log-mortality Ņ 4 ဖု ထု -10 20 40 60 0 80 Age







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All solved using Bayesian Hierarchical Models! (See *Demographic Forecasting* and the *YourCast* package.